

SUPPLEMENT TO “EXTREME POINTS AND MAJORIZATION: ECONOMIC APPLICATIONS”

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ANDREAS KLEINER

Department of Economics, Arizona State University

BENNY MOLDOVANU

Department of Economics, University of Bonn

PHILIPP STRACK

Department of Economics, Yale University

S.1. Schur-Convex Functions and Functionals

CONSIDER X_F AND X_G TO BE UNIFORM, discrete random variables, each taking n values $x_F = (x_F^1, \dots, x_F^n)$ and $x_G = (x_G^1, \dots, x_G^n)$, respectively. Then

$$x_F \prec_{\text{dm}} x_G \Leftrightarrow F^{-1} \prec G^{-1} \Leftrightarrow G \prec F,$$

where \prec_{dm} denotes the classical discrete majorization relation due to Hardy, Littlewood, and Polya. Thus, discrete majorization is equivalent to the present majorization relation applied to quantile functions. A function $V : \mathbb{R}^n \rightarrow \mathbb{R}$ is *Schur-convex* (*concave*) if $V(\mathbf{x}) \geq V(\mathbf{y})$ ($V(\mathbf{x}) \leq V(\mathbf{y})$) whenever $\mathbf{x} \succ_{\text{dm}} \mathbf{y}$. If V is a symmetric function, and if all its partial derivatives exist, then the *Schur–Ostrovski criterion* says that V is *Schur-convex* (*concave*) if and only if

$$(x_i - x_j) \left(\frac{\partial V}{\partial x_i} - \frac{\partial V}{\partial x_j} \right) \geq (\leq) 0 \quad \text{for all } x.$$

It is useful to have a similar characterization for continuous majorization. Chan, Proschan, and Sethuraman (1987) showed that a law-invariant,¹ Gâteaux-differentiable functional $V : L^1(0, 1) \rightarrow \mathbb{R}$ respects the majorization relation on $L^1(0, 1)$, if and only if its *Gâteaux-derivatives* in specially defined directions are non-positive. The considered directions are of the form

$$h = \lambda_1 \mathbf{1}_{(a,b)} + \lambda_2 \mathbf{1}_{(c,d)}$$

with $0 \leq a < b < c < d \leq 1$ and $\lambda_1 \geq 0 \geq \lambda_2$ such that $\lambda_1(b - a) + \lambda_2(d - c) = 0$. Note that the function h takes at most two values that are different from zero, and is decreasing on $[a, b] \cup [c, d]$. Moreover, $\int_0^1 h(t) dt = 0$.

This result also yields a simple intuition for the Fan–Lorentz theorem in the case where K is differentiable. Consider a monotonic f and note that, for any direction h ,

Andreas Kleiner: andykleiner@gmail.com

Benny Moldovanu: mold@uni-bonn.de

Philipp Strack: philipp.strack@gmail.com

¹This means that the functional is constant over the equivalence class of functions with the same non-decreasing rearrangement. This replaces the symmetry in the discrete formulation.

the Gâteaux-derivative of the functional $V(f) = \int_0^1 K(f(t), t) dt$ is given by

$$\delta V(f, h) = \frac{d}{d\varepsilon} \int_0^1 K(f(t) + \varepsilon h(t), t) dt \Big|_{\varepsilon=0} = \int_0^1 K_f(f(t), t) h(t) dt,$$

where the last equality follows by interchanging the order of differentiation and integration.² The Fan–Lorentz conditions imply together that

$$\frac{dK_f}{dt} = f_t \cdot K_{ff} + K_{ft} \geq 0.$$

For a direction h such that $\int_0^1 h(t) dt = 0$, and such that h is a decreasing two-step function as defined above, we obtain that

$$\delta V(f, h) = \int_0^1 K_f(f(t), t) h(t) dt \leq 0.$$

Hence, the Fan–Lorentz functional $V(f) = \int_0^1 K(f(t), t) dt$ is Schur-concave by the result of Chan, Proschan, and Sethuraman (1987).

S.2. Decision-Making Under Uncertainty

We briefly illustrate here how our insights can be applied in order to understand how agents with non-expected utility preferences choose among risky prospects.

S.2.1. Rank-Dependent Utility and Choquet Capacities

Quiggin (1982) and Yaari (1987) axiomatically derived utility functionals with rank-dependent assessments of probabilities of the form³

$$U(F) = \int_0^1 v(t) d(g \circ F)(t),$$

where F is the distribution of a random variable on the interval $[0, 1]$, $v : [0, 1] \rightarrow \mathbb{R}$ is continuous, strictly increasing, and bounded, and where $g : [0, 1] \rightarrow [0, 1]$ is strictly increasing, continuous, and onto. The function v represents a transformation of monetary payoffs, while the function g represents a transformation of probabilities.⁴

The case $g(x) = x$ yields the classical von Neumann–Morgenstern expected utility model where risk aversion is equivalent to v being concave. The case $v(x) = x$ yields Yaari’s (1987) dual utility theory, where risk aversion is equivalent to g being concave. Because of the possible interactions between v and g , it is not clear what properties yield risk aversion in the general rank-dependent model. Using integration by parts, we can

²This is allowed since K is convex in f .

³Their theory is a bit more general (e.g., it allows a more general domain for the functions v and F). We keep here a framework that is compatible with the rest of the paper.

⁴For the sake of brevity, we assume below that both g and v are twice differentiable. Since the Fan–Lorentz result does not require differentiability, the observations below generalize.

also write

$$\begin{aligned} U(F) &= \int_0^1 v(t) d(g \circ F)(t) = v(1) - \int_0^1 v'(t)(g \circ F)(t) dt \\ &= v(1) + \int_0^1 K(F(t), t) dt, \end{aligned}$$

where

$$K(F, t) = -v'(t)(g \circ F),$$

and where we used $g(0) = 0$ and $g(1) = 1$. Then

$$\frac{\partial^2 K(F, t)}{\partial F \partial t} = -g'(F(t))v''(t) \geq 0$$

for all t if and only if v is concave. Similarly,

$$\frac{\partial^2 K(F, t)}{\partial^2 F} = -g''(F(t))v'(t) \geq 0$$

for all t if and only if g is concave.

Hence, the Fan–Lorentz conditions are satisfied if and only if $v'' \leq 0$ and $g'' \leq 0$. As a consequence, the utility functional $U = \int_0^1 v(t) d(g \circ F)(t)$ is Schur-concave, and the agent whose preferences are represented by U is *risk averse*, exactly as under standard expected utility.⁵

Another important strand of the literature on non-expected utility considers ambiguity aversion. The main tool is the *Choquet integral* with respect to a (convex) *capacity* (this is unrelated to the Choquet representation used above!). Analogously to the derivations above, it can be shown that the Choquet integral yields a Schur-concave functional if and only if it is computed with respect to a convex capacity.

S.2.2. A Portfolio Choice Problem

Dybvig (1988) studied a simplified version of the following problem:

$$\begin{aligned} \min_X \mathbb{E}[XY] \\ \text{s.t. } X \geq_{cv} Z, \end{aligned}$$

where Y and Z are given random variables. Y represents here the distribution of a pricing function over the states of the world, and the goal is to choose, given Y , the cheapest contingent claim X that is less risky than a given claim Z . To make the problem well-defined, Y needs to be essentially bounded and X, Z must be integrable. Recalling that

$$X \geq_{cv} Z \Leftrightarrow F_X \succ F_Z \Leftrightarrow F_X^{-1} \prec F_Z^{-1},$$

⁵The equivalence between the concavity of the functions v and g and risk aversion has been pointed out by Hong, Karni, and Safra (1987), who built on Machina (1982).

we obtain that

$$\mathbb{E}[XY] \geq \int_0^1 F_Y^{-1}(1-t)F_X^{-1}(t) dt \geq \int_0^1 F_Y^{-1}(1-t)F_Z^{-1}(t) dt,$$

where the first inequality follows by the rearrangement inequality of [Hardy, Littlewood, and Polya \(1929\)](#) (the anti-assortative part!), and where the second inequality follows by the Fan–Lorentz theorem.

By choosing a random variable X that has the same distribution as Z and that is anti-comonotonic with Y ,⁶ the lower bound $\int_0^1 F_Y^{-1}(1-t)F_Z^{-1}(t) dt$ is attained, and hence such a choice solves the portfolio choice problem.⁷

If $Y' \leq_{cv} Y$, we obtain by the Fan–Lorentz inequality (now applied to the functional with argument F_Y^{-1}) that

$$\sup_{X \succ_{cv} Z} \mathbb{E}[XY] = \int_0^1 F_Y^{-1}(1-t)F_Z^{-1}(t) dt \geq \int_0^1 F_{Y'}^{-1}(1-t)F_Z^{-1}(t) dt = \sup_{X \succ_{cv} Z} \mathbb{E}[XY'].$$

In other words, a decision maker that becomes more informed (in the Blackwell sense) will bear a lower cost.

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⁶This can always be done if the underlying probability space is non-atomic. A random vector (X, Y) is anti-comonotonic if there exists a random variable W and non-decreasing functions h_1, h_2 such that $(X, Y) \stackrel{\text{dist}}{=} (h_1(W), -h_2(W))$.

⁷For more details on this problem, see [Dana \(2005\)](#) and the literature cited there. It does not use the Fan–Lorentz inequality.